STAT 423/523 Statistical Methods for Engineers and Scientists

Lecture 9: Multifactor Analysis of Variance II

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Now we consider the two-way ANOVA with replication. That is for each combination of the two factors, we have more than one observation.

$$X_{ijk} = \mu_{ij} + \epsilon_{ijk},$$

where the index ijk denotes the kth observation in the ith level of factor A and the jth level of factor B.

For simplicity, we assume $i=1,2,\ldots,I$, $j=1,2,\ldots,J$, and $k=1,2,\ldots,K$ — equal number of observations in each cell.

We rewrite μ_{ij} as in the additive model but with an additional term for the interaction effect:

$$\mu_{ij} = \mu + \alpha_i + \beta_j + \gamma_{ij}.$$

- ▶ The term γ_{ij} is called the **interaction effect**.
- ▶ We don't include the interaction term in two-way ANOVA without replication because we can't estimate it (more parameters than observations).
- ▶ With replication, we can estimate the interaction effect.
- The number of parameters for μ_{ij} is IJ. So we need to show that the number of above model is IJ as well.

$$\mu_{ij} = \mu + \alpha_i + \beta_j + \gamma_{ij}.$$

- ▶ The number of parameters is 1 + I + J + IJ = (I + 1)(J + 1).
- ► The constraints are:

$$\sum_i \alpha_i = 0,$$
 $\sum_j \beta_j = 0,$ $\sum_j \gamma_{ij} = 0$ for all $j = 1, \dots, J,$ $\sum_j \gamma_{ij} = 0$ for all $i = 1, \dots, I.$

The number of constraints is I + J + 2.

The number of **independent** constraints is I + J + 1.

► Therefore, the effective number of parameters is (I+1)(J+1) - (I+J+1) = IJ.

$$\mu_{ij} = \mu + \alpha_i + \beta_j + \gamma_{ij}.$$

- $ightharpoonup \alpha_i$: the (main) effect of the *i*th level of factor A.
- \triangleright β_j : the (main) effect of the *j*th level of factor B.
- $ightharpoonup \gamma_{ij}$: the interaction effect between the *i*th level of factor A and the *j*th level of factor B.

Sample Means

Similar to the one-way ANOVA, we can calculate the sample means for each level of the two factors and the interaction effect:

$$\bar{X}_{i..} = \frac{1}{JK} \sum_{j=1}^{J} \sum_{k=1}^{K} X_{ijk}, \qquad \bar{X}_{.j.} = \frac{1}{IK} \sum_{i=1}^{I} \sum_{k=1}^{K} X_{ijk}$$

$$\bar{X}_{ij.} = \frac{1}{K} \sum_{k=1}^{K} X_{ijk}, \qquad \bar{X}_{...} = \frac{1}{IJK} \sum_{i=1}^{I} \sum_{j=1}^{K} \sum_{k=1}^{K} X_{ijk}.$$

- $ightharpoonup \bar{X}_{i}$: the sample mean for the *i*th level of factor A.
- $lackbox{} \bar{X}_{\cdot j}$: the sample mean for the jth level of factor B.
- $ightharpoonup \bar{X}_{ij}$: the sample mean for the ijth cell.
- $ightharpoonup ar{X}$...: the grand mean.

Estimation

$$X_{ijk} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijk}.$$

Using method of moments, the estimators should satisfy

$$\bar{X}_{ij} = \hat{\mu} + \hat{\alpha}_i + \hat{\beta}_j + \hat{\gamma}_{ij} \quad (*)$$

- Averaging (*) over all i and j gives: $\bar{X}_{...} = \hat{\mu}$.
- Averaging (*) over all i and fixing j gives: $\bar{X}_{.j.} = \hat{\mu} + \hat{\beta}_j \Longrightarrow \hat{\beta}_j = \bar{X}_{.j.} \bar{X}_{...}$
- ightharpoonup Similarly, $\hat{\alpha}_i = \bar{X}_{i\cdots} \bar{X}_{\cdots}$
- ▶ By plugging in the above estimators, we can solve for $\hat{\gamma}_{ij} = \bar{X}_{ij} \bar{X}_{i\cdots} \bar{X}_{\cdot j} + \bar{X}_{\cdots}$

Estimation

$$X_{ijk} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijk}.$$

The estimators are:

$$\begin{split} \hat{\mu} &= \bar{X}..., \\ \hat{\alpha}_i &= \bar{X}_{i..} - \bar{X}..., \\ \hat{\beta}_j &= \bar{X}_{.j.} - \bar{X}..., \\ \hat{\gamma}_{ij} &= \bar{X}_{ij.} - \bar{X}_{i...} - \bar{X}_{.j.} + \bar{X}... \\ \hat{\epsilon}_{ijk} &= X_{ijk} - \bar{X}_{ij.}. \end{split}$$

Sum of Squares

$$X_{ijk} = \hat{\mu} + \hat{\alpha}_i + \hat{\beta}_j + \hat{\gamma}_{ij} + \hat{\epsilon}_{ijk}.$$

Now we can define the sum of squares for each term in the model:

$$SSE = \sum_{i} \sum_{j} \sum_{k} \hat{\epsilon}_{ijk}^{2} = \sum_{i} \sum_{j} \sum_{k} (X_{ijk} - \bar{X}_{ij}.)^{2}, \qquad df = IJ(K - 1)$$

$$SSA = \sum_{i} \sum_{j} \sum_{k} \hat{\alpha}_{i}^{2} = JK \sum_{i} (\bar{X}_{i..} - \bar{X}_{...})^{2}, \qquad df = I - 1$$

$$SSB = \sum_{i} \sum_{j} \sum_{k} \hat{\beta}_{j}^{2} = IK \sum_{j} (\bar{X}_{.j}. - \bar{X}_{...})^{2}, \qquad df = J - 1$$

$$SSAB = \sum_{i} \sum_{j} \sum_{k} \hat{\gamma}_{ij}^{2} = K \sum_{i} \sum_{j} (\bar{X}_{ij}. - \bar{X}_{i..} - \bar{X}_{.j}. + \bar{X}_{...})^{2}, \qquad df = (I - 1)(J - 1)$$

$$SST = \sum_{i} \sum_{j} \sum_{k} (X_{ijk} - \bar{X}_{...})^{2}. \qquad df = IJK - 1$$

Also, we have SST = SSA + SSB + SSAB + SSE.

Mean Squares

We can define the mean squares and give their expected values:

$$MSE = \frac{SSE}{IJ(K-1)}$$

$$MSA = \frac{SSA}{I-1}$$

$$E(MSE) = \sigma^{2},$$

$$E(MSA) = \sigma^{2} + \frac{JK}{I-1} \sum_{i} \alpha_{i}^{2},$$

$$MSB = \frac{SSB}{J-1}$$

$$E(MSB) = \sigma^{2} + \frac{IK}{J-1} \sum_{j} \beta_{j}^{2},$$

$$MSAB = \frac{SSAB}{(I-1)(J-1)}$$

$$E(MSAB) = \sigma^{2} + \frac{K}{(I-1)(J-1)} \sum_{i} \sum_{j} \gamma_{ij}^{2}.$$

To test the main effect of factor A, we consider the following hypotheses:

$$H_0: \alpha_1 = \alpha_2 = \cdots = \alpha_I = 0,$$

 $H_1:$ at least one $\alpha_i \neq 0.$

Under null the model is

$$X_{ijk} = \mu + \beta_j + \gamma_{ij} + \epsilon_{ijk}$$

Under alternative the model is

$$X_{ijk} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijk}$$

We summarize the two nested models as

Model	SS Model	SS Error	Difference in SS Error	Difference in df
Full	SSA+SSB+SSAB	SSE		
Reduced	SSB + SSAB	SSE+SSA	SSA	I-1

Model	SS Model	SS Error	Difference in SS Error	Difference in df
Full	SSA+SSB+SSAB	SSE		
Reduced	SSB + SSAB	SSE+SSA	SSA	l-1

Therefore, the F-test should be

$$F_A = \frac{SSA/(I-1)}{SSE/[IJ(K-1)]} = \frac{MSA}{MSE} \sim F_{I-1,IJ(K-1)}(\text{under null})$$

We should reject null when $F_A > F_{\alpha,I-1,IJ(K-1)}$.

Similarly, to test the main effect of factor B, we consider the following hypotheses:

$$H_0: \beta_1 = \beta_2 = \cdots = \beta_J = 0,$$

 $H_1:$ at least one $\beta_j \neq 0.$

The F-test should be

$$F_B = rac{MSB}{MSE} \sim F_{J-1,IJ(K-1)} (ext{under null})$$

We should reject null when $F_B > F_{\alpha,J-1,IJ(K-1)}$.

To test the interaction effect, we consider the following hypotheses:

$$H_0: \gamma_{11} = \gamma_{12} = \cdots = \gamma_{IJ} = 0,$$

 $H_1:$ at least one $\gamma_{ij} \neq 0.$

The F-test should be

$$F_{AB} = \frac{MSAB}{MSE} \sim F_{(I-1)(J-1),IJ(K-1)}(\text{under null})$$

We should reject null when $F_{AB} > F_{\alpha,(I-1)(J-1),IJ(K-1)}$.

Questions: What are the full model and reduced model in this case?

ANOVA Table

The ANOVA table for the two-way ANOVA with replication is:

Source	SS	df	MS	F
Factor A	SSA	I-1	MSA	F_A
Factor B	SSB	J-1	MSB	F_B
Interaction	SSAB	(I-1)(J-1)	MSAB	F_{AB}
Error	SSE	IJ(K-1)	MSE	
Total	SST	IJK-1		

Example

Data: thermal properties of asphalt mix under three different binder grades and three different coarse aggregate contents.

	Coarse Aggregate Content (%)			
Asphalt Binder Grade	38	41	44	$\overline{x}_{i\cdots}$
PG58	.835, .845	.822, .826	.785, .795	.8180
PG64	.855, .865	.832, .836	.790, .800	.8297
PG70	.815, .825	.800, .820	.770, .790	.8033
$\overline{\overline{x}}_{.j.}$.8400	.8227	.7883	

Example

The ANOVA table is:

Source	DF	SS	MS	f	P
AsphGr	2	.0020893	.0010447	14.12	0.002
AggCont	2	.0082973	.0041487	56.06	0.000
Interaction	4	.0003253	.0000813	1.10	0.414
Error	9	.0006660	.0000740		
Total	17	.0113780			

In some experiments, one or more factors are randomly selected from a larger population.

- ▶ The factors that are pre-selected are called **fixed effects**.
- The factors that are randomly selected are called random effects.
- If a multi-factor model contains both fixed and random effects, it is called a mixed effect model.

Two-way ANOVA model with two fixed effects (with replication):

$$X_{ijk} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijk}.$$

 $ightharpoonup \alpha_i$ and β_j 's are unknown fixed parameters.

Now we let the second factor to be random:

$$X_{ijk} = \mu + \alpha_i + B_j + G_{ij} + \epsilon_{ijk}.$$

- $ightharpoonup B_j$ is the random effect of the jth level of factor B.
- $ightharpoonup G_{ij}$ is the interaction effect between the ith level of factor A and the jth level of factor B.

$$X_{ijk} = \mu + \alpha_i + B_j + G_{ij} + \epsilon_{ijk}.$$

Certain assumptions are made for the random effects:

- $B_j \sim N(0, \sigma_B^2).$
- $G_{ij} \sim N(0, \sigma_G^2).$

$$X_{ijk} = \mu + \alpha_i + B_j + G_{ij} + \epsilon_{ijk}.$$

Now we consider the constraints:

- $\sum_{i} \alpha_{i} = 0$ because it's on the fixed effect, so we keep it.
- $ightharpoonup \sum_{j} B_{j} = 0$ we should **not** assume it because of random sampling.
- $ightharpoonup \sum_{i} G_{ij} = 0$ we should **not** assume it as well.
- $ightharpoonup \sum_i G_{ij} = 0$ it depends.

$$X_{ijk} = \mu + \alpha_i + B_j + G_{ij} + \epsilon_{ijk}.$$

The unrestricted model assumes:

- $ightharpoonup \epsilon_{ijk} \sim N(0, \sigma^2).$
- $\triangleright B_j \sim N(0, \sigma_B^2).$
- $ightharpoonup G_{ij} \sim N(0, \sigma_G^2).$
- $ightharpoonup \sum_i \alpha_i = 0.$
- ▶ B_j and G_{ij} are independent of each other.

$$X_{ijk} = \mu + \alpha_i + B_j + G_{ij} + \epsilon_{ijk}.$$

The restricted model assumes:

- $B_j \sim N(0, \sigma_B^2).$
- $\triangleright \sum_i \alpha_i = 0.$
- $\triangleright \sum_{i} G_{ij} = 0.$
- Now G_{ij} 's are correlated.

The direct consequences of the different assumptions are the expected MS values.

Unrestricted Model

$$E(MSE) = \sigma^2$$

$$E(MSA) = \sigma^2 + K\sigma_G^2 + \frac{JK}{I-1} \sum_i \alpha_i^2$$

$$E(MSB) = \sigma^2 + K\sigma_G^2 + IK\sigma_B^2$$

$$E(MSAB) = \sigma^2 + K\sigma_G^2$$

Restricted Model

$$E(MSE) = \sigma^2$$

 $E(MSA) = \sigma^2 + K\sigma_G^2 + \frac{JK}{I-1} \sum_i \alpha_i^2$
 $E(MSB) = \sigma^2 + IK\sigma_B^2$
 $E(MSAB) = \sigma^2 + K\sigma_G^2$

The further consequence is that we should use different F-values for the tests.

Test (H_0)	Unrestricted Model	Restricted Model
$\alpha_1 = \alpha_2 = \dots = \alpha_I = 0$ $\sigma_B^2 = 0$ $\sigma_G^2 = 0$		F = MSB/MSE